

# Policy learning under the $f$ -sensitivity model

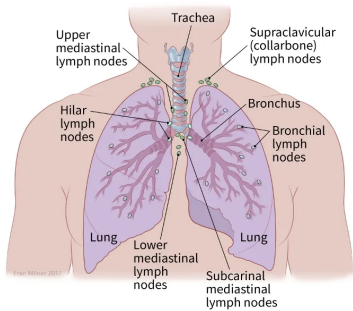
Zhimei Ren

University of Pennsylvania, Department of Statistics and Data Science

*IMSI, April 2026*

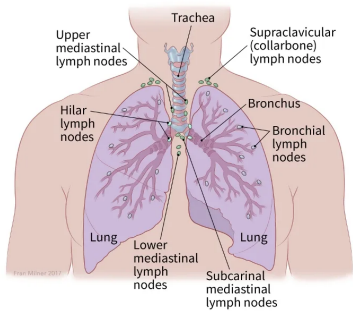
Joint work with Jingyuan Wang (NYU), Ruohan Zhan (UCL), Xinyu Liang (INSEAD), and Zhengyuan Zhou (NYU)

# Non-small cell lung cancer treatment



[Source: American Cancer Society]

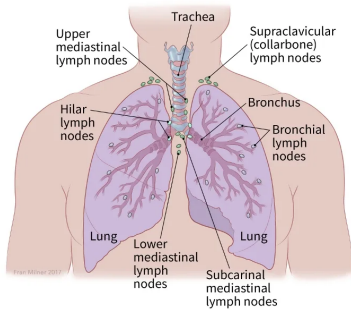
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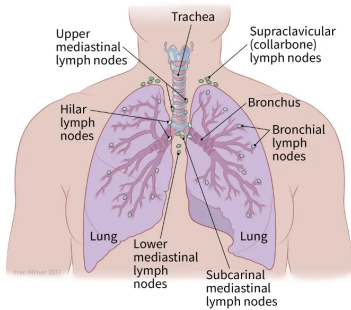


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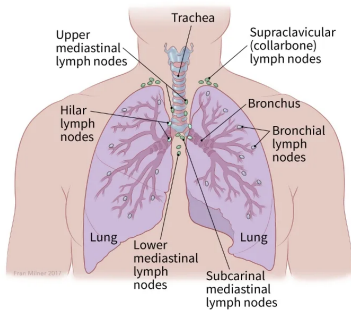


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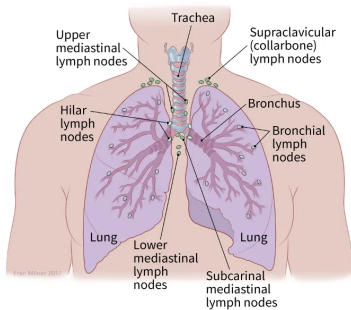


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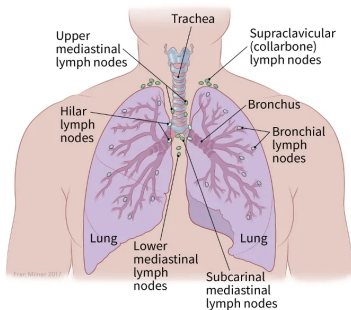


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## Treating stage IIIA NSCLC

The initial treatment for stage IIIA NSCLC may include some combination of [radiation therapy](#), [chemotherapy](#) (chemo), immunotherapy, and/or [surgery](#). For this reason, planning treatment for stage IIIA NSCLC often requires input from a medical oncologist, radiation oncologist, and a thoracic surgeon. Your treatment options depend on the size of the tumor, where it is in your lung, which lymph nodes it has spread to, your overall health, and how well you are tolerating treatment.

# Treatment for stage III NSCLC

Can we learn an individualized treatment policy for stage III NSCLC from observational data?

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<sup>1</sup>curated from Truveta's real-world electronic health record (EHR) database

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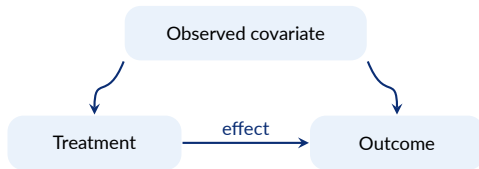
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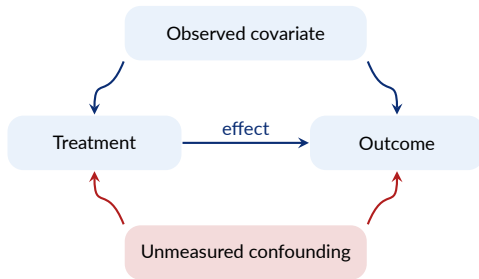
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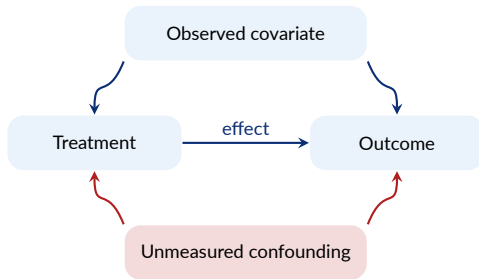


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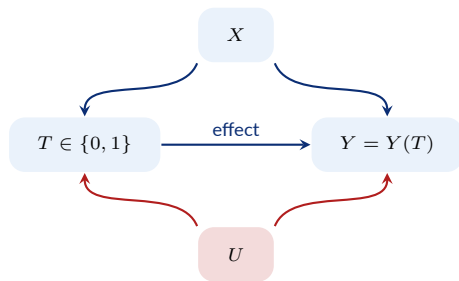
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Common in observational studies!

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# Characterizing unmeasured confounding

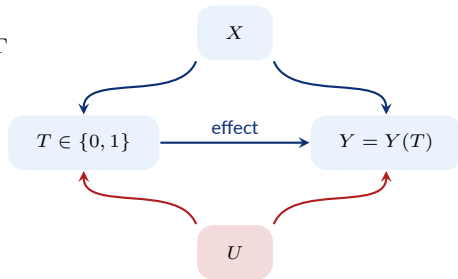


# Characterizing unmeasured confounding

- $\Gamma$ -sensitivity model [Rosenbaum (1987)]

$$\frac{1}{\Gamma} \leq \frac{\mathbb{P}(T = 1|X = x, U = u)}{\mathbb{P}(T = 0|X = x, U = u)} \bigg/ \frac{\mathbb{P}(T = 1|X = x, U = u')}{\mathbb{P}(T = 0|X = x, U = u')} \leq \Gamma$$

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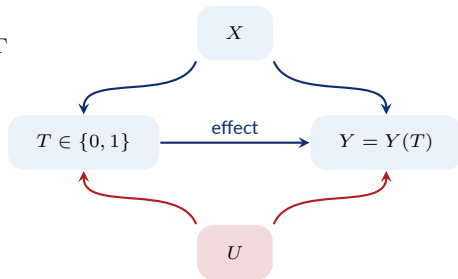
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- marginal  $\Gamma$ -sensitivity model [Tan, 2006]

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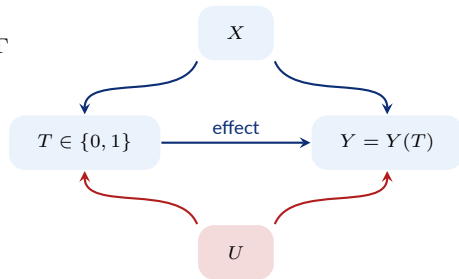
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$\Gamma \geq 1$  quantifies the strength of unmeasured confounding



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Assuming the (marginal)  $\Gamma$ -sensitivity model will be overly conservative if the case of extreme confounding is rare

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the true law of  $(Y(1), Y(0), X)$  is not point identifiable from the observed data

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the true law of  $(Y(1), Y(0), X)$  is **is partially identifiable** from the observed data **under sensitivity models**

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# Partial identification under the f-sensitivity model

Lemma. [Jin, R., Zhou '26]

Under the f-sensitivity model with parameter  $\rho$ , there is

$$D_f\left(\underbrace{P_{Y(t)|X=x,T=t}}_{\text{observable}} \parallel P_{Y(t)|X=x}\right) \leq \rho, \quad t = 0, 1$$

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$$P_X = P_X^{\text{obs}}, \quad D_f(P_{Y(t)|X=x,T=t}^{\text{obs}} \parallel P_{Y(t)|X=x}) \leq \rho, \quad \forall t = 0, 1$$

# Offline data collection mechanism

How is  $\mathcal{D} = \{X_i, T_i, Y_i\}_{i=1}^n$  collected? Consider two types of data collection mechanisms:

## I.i.d. data

- ▶  $(U_i, X_i, T_i, Y_i(0), Y_i(1)) \stackrel{\text{iid}}{\sim} P, i = 1, \dots, n$
- ▶ Observe  $(X_i, T_i, Y_i)$  with  $Y_i = Y_i(T_i)$
- ▶ Propensity score  $e(X) = \mathbb{P}(T = 1|X) \in [\epsilon, 1 - \epsilon]$

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# Distributionally robust policy learning

Learning objective:

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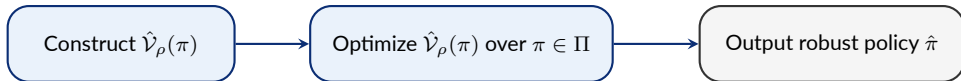
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# Policy evaluation: from primal to dual

Fix  $\pi \in \Pi$

- Need to estimate the policy value under the worst-case law of  $Y(1), Y(0) \mid X$

$$\begin{aligned}\mathcal{V}_\rho(\pi) &= \min_{Q \in \mathcal{Q}_f(\rho)} \mathbb{E}_Q[Y(\pi(X))] \\ &= \mathbb{E}_{P_X} \left[ \pi(X) \cdot \min_{Q \in \mathcal{Q}_f(\rho)} \mathbb{E}_Q[Y(1)|X] + (1 - \pi(X)) \cdot \min_{Q \in \mathcal{Q}_f(\rho)} \mathbb{E}_Q[Y(0)|X] \right]\end{aligned}$$

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Lemma. [Jin, R., Zhou '26]

Under the  $f$ -sensitivity model with parameter  $\rho$ , for any  $t = 0, 1$ ,

$$\min_{Q \in \mathcal{Q}_f(\rho)} \mathbb{E}_Q[Y(t)|X] = - \min_{\alpha \geq 0, \eta \in \mathbb{R}} \mathbb{E}_P \left[ \alpha f^* \left( - \frac{Y(t) + \eta}{\alpha} \right) + \eta + \alpha \rho \mid X, T = t \right],$$

where  $f^*(z) = \sup_{x \geq 0} \{zx - f(x)\}$  is the convex conjugate of  $f$ .

# Policy evaluation: estimator

Dual opt. problem:  $\min_{\alpha \geq 0, \eta \in \mathbb{R}} \mathbb{E}_P \left[ \underbrace{\alpha f^* \left( -\frac{Y(t) + \eta}{\alpha} \right) + \eta + \alpha \rho}_{R_t(\alpha, \rho)} \mid X, T = t \right]$

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► Plug-in approach?

$$\hat{\mathcal{V}}_\rho(\pi) = \frac{1}{n} \sum_{i=1}^n \left[ \frac{\mathbb{1}\{T_i = \pi(X_i)\}}{\hat{e}(X_i)} \cdot R_{\pi(X_i)}(\hat{\alpha}(X_i), \hat{\eta}(X_i)) \right]$$

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- Nuisance functions  $\hat{e}(\cdot), \hat{\alpha}(\cdot), \hat{\eta}(\cdot)$  are estimated by the same data  $\mathcal{D} \Rightarrow$  inefficient!

# Policy evaluation: cross-fitting estimator

**Fold 1**

Solve dual opt. obtain  $\hat{\alpha}_t^{(1)}(x), \hat{\eta}_t^{(1)}(x)$

**Fold 2**

Estimate  $e(x)$  and  $r_t := \mathbb{E}[R_t(\hat{\alpha}^{(1)}, \hat{\eta}^{(1)}) | X] \rightsquigarrow (\hat{e}^{(1)}, \hat{r}_t^{(1)})$

**Fold 3**

$$\hat{\mathcal{V}}_{\rho}^{(1)}(\pi) = \sum_{i \in \mathcal{D}_3} \frac{\mathbb{1}\{T_i = \pi(X_i)\}}{\hat{e}^{(1)}(\pi(X_i))} \left( R_{\pi(X_i)}(\hat{\alpha}^{(1)}, \hat{\eta}^{(1)}) - \hat{r}^{(1)}(X_i) \right) + \hat{r}^{(1)}(X_i)$$

# Policy evaluation: cross-fitting estimator

**Fold 1**

$$\hat{\mathcal{V}}_{\rho}^{(2)}(\pi) = \sum_{i \in \mathcal{D}_1} \frac{\mathbb{1}\{T_i = \pi(X_i)\}}{\hat{e}^{(2)}(\pi(X_i))} \left( R_{\pi(X_i)}(\hat{\alpha}^{(2)}, \hat{\eta}^{(2)}) - \hat{r}^{(2)}(X_i) \right) + \hat{r}^{(2)}(X_i)$$

**Fold 2**

Solve dual opt. obtain  $\hat{\alpha}_t^{(2)}(x), \hat{\eta}_t^{(2)}(x)$

**Fold 3**

Estimate  $e(x)$  and  $r_t := \mathbb{E}[R_t(\hat{\alpha}^{(2)}, \hat{\eta}^{(2)}) | X] \rightsquigarrow (\hat{e}^{(2)}, \hat{r}_t^{(2)})$

# Policy evaluation: cross-fitting estimator

**Fold 1**

Estimate  $e(x)$  and  $r_t := \mathbb{E}[R_t(\hat{\alpha}^{(3)}, \hat{\eta}^{(3)}) | X] \rightsquigarrow (\hat{e}^{(3)}, \hat{r}_t^{(3)})$

**Fold 2**

$$\hat{\mathcal{V}}_{\rho}^{(3)}(\pi) = \sum_{i \in \mathcal{D}_2} \frac{\mathbb{1}\{T_i = \pi(X_i)\}}{\hat{e}^{(3)}(\pi(X_i))} \left( R_{\pi(X_i)}(\hat{\alpha}^{(3)}, \hat{\eta}^{(3)}) - \hat{r}^{(3)}(X_i) \right) + \hat{r}^{(3)}(X_i)$$

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Solve dual opt. obtain  $\hat{\alpha}_t^{(3)}(x), \hat{\eta}_t^{(3)}(x)$



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Theorem (Double robustness of the policy value estimator). [Wang, R., Zhan, Liang, Zhou '26+]

Under mild regularity conditions, suppose for each fold  $k \in [3]$

$$\begin{aligned} \|\hat{e}^{(k)} - e\|_2 &= o_p(n^{-\gamma_1}), \quad \|\hat{r}_t^{(k)} - r_t\|_2 = o_p(n^{-\gamma_2}), \quad t = 0, 1 \\ \|(\hat{\alpha}^{(k)}, \hat{\eta}^{(k)}) - (\alpha, \eta)\|_2 &= o_p(n^{-1/4}), \quad \|(\hat{\alpha}^{(k)}, \hat{\eta}^{(k)}) - (\alpha, \eta)\|_{\infty} = o_p(1), \end{aligned}$$

for some  $\gamma_1, \gamma_2 > 0$  satisfying  $\gamma_1 + \gamma_2 > 1/2$ . Then for any  $\pi$ ,

$$\sqrt{n}(\hat{\mathcal{V}}_{\rho}(\pi) - \mathcal{V}_{\rho}(\pi)) \Rightarrow \mathcal{N}(0, \sigma_{\pi}^2).$$

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- ▶ Efficient implementation for tree-based policy class via dynamic programming

# Policy optimization: regret upper bound

Theorem (Regret upper bound). [Wang, R., Zhan, Liang, Zhou '26+]

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►  $\kappa(\Pi)$ : entropy integral of  $\Pi$  under the Hamming distance  $\rightsquigarrow$  complexity of  $\Pi$

# Policy optimization: regret lower bound

Theorem (Minimax lower bound). [Wang, R., Zhan, Liang, Zhou '26+]

Suppose  $\rho \leq 0.2$  and  $n \geq \text{VCdim}(\Pi)^2$ . There exists a distribution over  $(X, U, T, Y(0), Y(1))$  satisfying the same condition in the upper bound result, such that for any  $\hat{\pi}$ ,

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When  $\text{Err}(e) \times \text{Err}(r_t) = O(n^{-1/2})$  and  $\text{Err}(\alpha) \vee \text{Err}(\eta) = O(n^{-1/4})$ ,

**Minimax optimal!**

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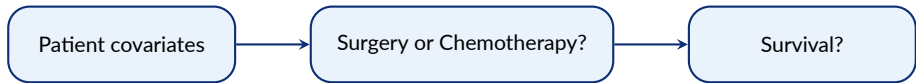
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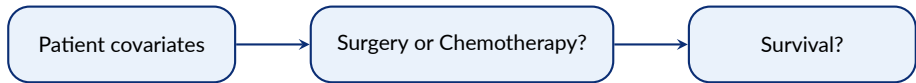
Similar algorithm (w/o cross-fitting)

$$\mathbb{E}[\text{regret}(\hat{\pi})] = O\left(\frac{\kappa(\Pi)}{n} \sqrt{\sum_i \frac{1}{\epsilon_i^2}}\right)$$

# NSCLC dataset revisited

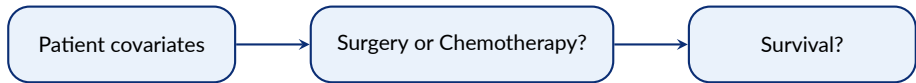


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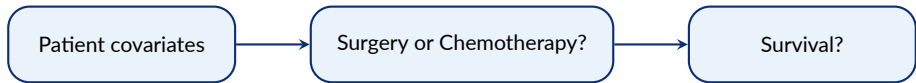
- $U$  =BMI or  $U$  =cholesterol level

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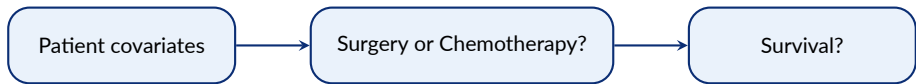
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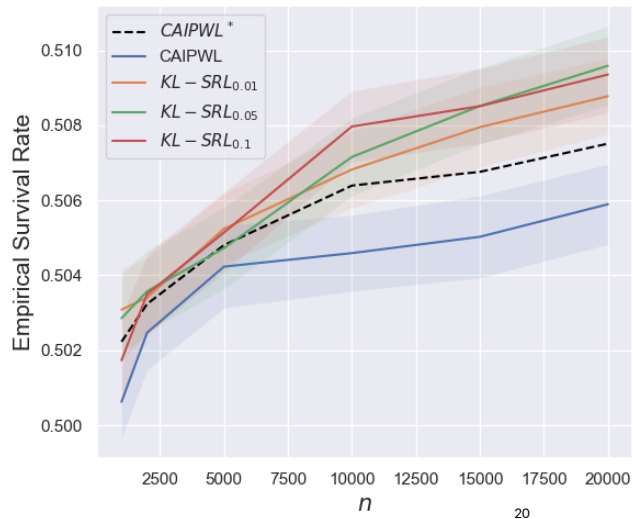
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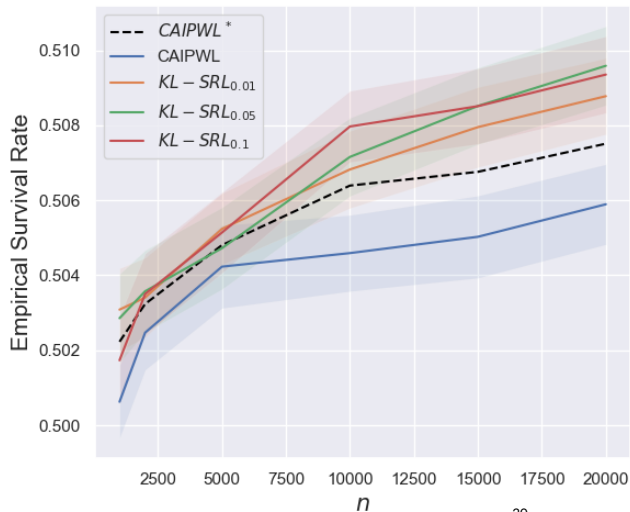


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- ▶ The other half  $\mathcal{D}_2$  w/o BMI or cholesterol level: learn a robust policy  $\hat{\pi}$  and evaluate it
- ▶ Choice of  $\rho$ ? Hard to estimate in general  $\Rightarrow$  conduct sensitivity analysis by varying  $\rho$

# BMI as unmeasured confounder

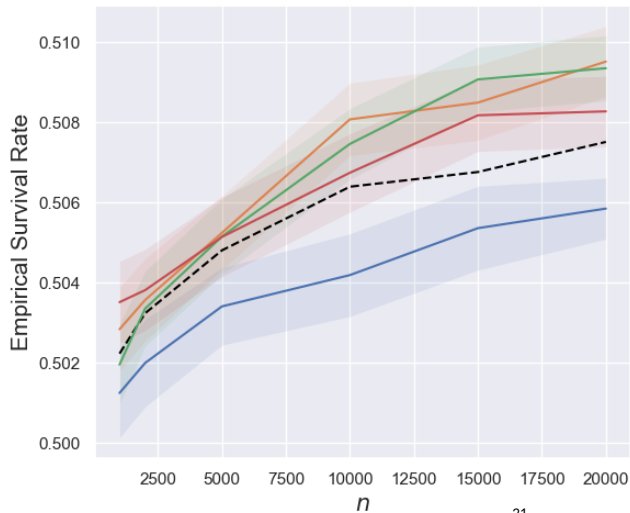


# BMI as unmeasured confounder

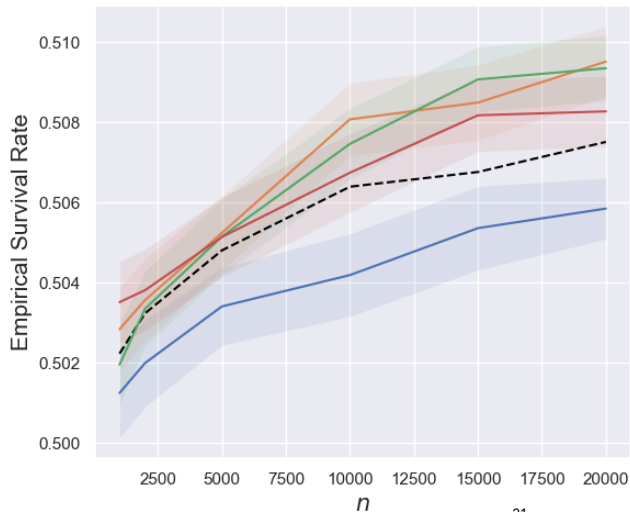


- ▶  $f(x) = x \log x \rightsquigarrow$  KL divergence
- ▶ Baseline:  $CAIPWL$   
[Zhou, Athey and Wager '22]
- ▶ Oracle baseline:  $CAIPWL$   
with BMI observed

# Cholesterol level as unmeasured confounder

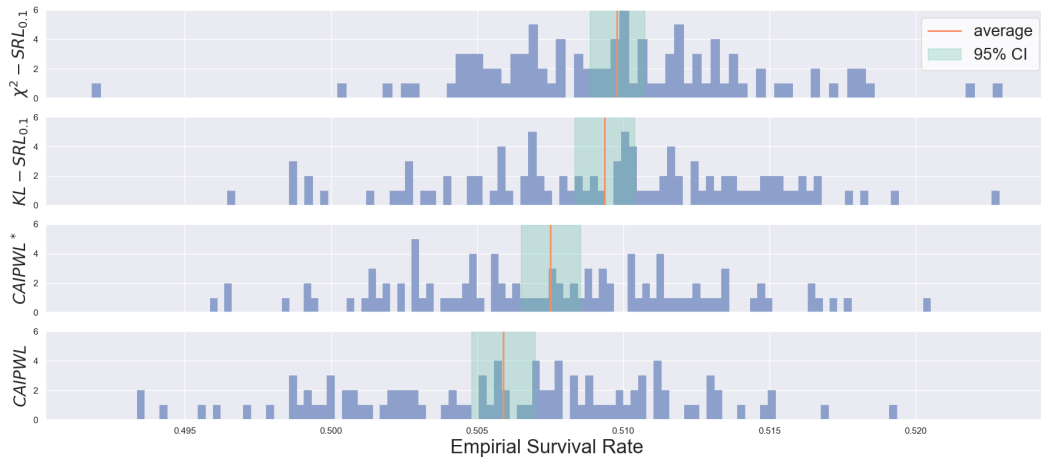


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# Histogram of survival rate



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- ▶ Efficient **policy evaluation** and **optimization** algorithms under the f-sensitivity model
- ▶ Applies to both i.i.d. and adaptively collected data
- ▶ Empirical results on NSCLC dataset suggest the robust policy can outperform the baseline

## Related work:

- ▶ Ying Jin, Zhimei Ren, and Zhengyuan Zhou. "Sensitivity Analysis Under the  $f$ -Sensitivity Model: A Distributional Robustness Perspective." *Operations Research* (2026).
- ▶ Jinyuan Wang, Zhimei Ren, Ruohan Zhan, and Zhengyuan Zhou. "Distributionally robust policy learning under concept drifts." *ICML* (2025).
- ▶ Jinyuan Wang, Zhimei Ren, Ruohan Zhan, Xinyu Liang and Zhengyuan Zhou. "Optimal policy learning under the  $f$ -sensitivity model." *in preparation* (2026+).

Thank you!